Sparse Regularization by Evolving the ℓ_1 Subgradient

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Sparse recovery

• Goal: recover a sparse vector $u \in \mathbb{R}^n$ from noisy measurements

$$b = Au + \omega.$$

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- Given A and b, we have **two tasks**:
 - 1. variable/predictor selection: find the support of u
 - 2. estimation: predict the values of u
- Largely many applications and several existing approaches

ℓ_1 subgradient

- Proposed method: variable selection based on ℓ_1 -subgradient p
- Subdifferential of convex function \boldsymbol{f}

$$\partial f(x) = \{ p : f(y) \ge f(x) + \langle p, y - x \rangle, \ \forall y \in \text{dom}f \}.$$

 $p \in \partial f(x)$ is a subgradient of f at x.

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- Subdifferential of $|\cdot|$:

$$\partial |x| = \begin{cases} \{1\}, & x > 0; \\ [-1,1], & x = 0; \\ \{-1\}, & x < 0. \end{cases}$$

 \implies given that $p \in \partial |x|$, then

$$x \begin{cases} \geq 0, & \text{if } p = 1; \\ = 0, & \text{if } p \in (-1, 1); \\ \leq 0, & \text{if } p = -1. \end{cases}$$

• ℓ_1 subdifferential:

$$\partial \|u\|_1 = \partial |u_1| \times \cdots \partial |u_n|.$$

 \implies given that $p\in\partial\|u\|_1,$ then

$$u_i \begin{cases} \ge 0, & \text{if } p_i = 1; \\ = 0, & \text{if } p_i \in (-1, 1); \\ \le 0, & \text{if } p_i = -1. \end{cases}$$

- $u_i = \pm 1 \implies u_i$ can be nonzero.
- we select predictors by computing *p*.

Sparse variable selection

• Suppose $p \in \partial \|u\|_1$

 $u \in \mathbb{R}^n$ is sparse \iff few $p_i = \pm 1$

- Assume A is short and wide (few rows and more columns)
- $p \in \partial \|u\|_1 \cap \mathcal{R}(A^T) \implies u$ tends to sparse
- Subgaussian random A of appropriate size \implies sparse u w.h.p.

Data fitting

We shall compute p such that

- sparsity: $p \in \partial \|u\|_1 \cap \mathcal{R}(A^T)$
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Proposed system:

$$\dot{p}_{+}(t) = A^{*}(b - Au(t)),$$
 (1a)
 $p(t) \in \partial \|u(t)\|_{1}.$ (1b)

Initial solution: p(0) = 0, u(0) = 0. Notation:

- $\dot{p}_+(t)$: right derivative of p(t)
- $A^* = \frac{1}{m}A^T$: normalized adjoint
- $\partial \| \cdot \|_1$: ℓ_1 subdifferential

Known as inverse-scale space (ISS) with total variation

Toy example

• Single real measurement

$$b = \mathbf{a}^T u + \epsilon \in \mathbb{R}$$

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• At time $t_1 = |ba_1|^{-1}$,

 \Rightarrow

 $p_1(t_1) = \operatorname{sign}(ba_1), \quad p_2(t_1), \dots, p_n(t_1) \in (-1, 1).$

Hence, $u_1(t_1)$ can be nonzero.

Under technical assumptions:

- p is right continuously differentiable, and
- *u* is right continuous,

 $u(t_1)$ must be the solution to

minimize
$$\|\mathbf{a}^T u - b\|_2^2$$
 s.t. $p_1(t_1) \cdot u_1 \ge 0, \ u_2 = \cdots = u_n = 0.$

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$$\implies u_{1}(t_{1}) = \frac{b}{a_{1}}, \quad u_{2}(t_{1}) = \dots = u_{n}(t_{1}) = 0.$$
Easy to verify

 $p(t_1) \in \partial \|u(t_1)\|_1.$

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For $t > t_1$, $p(t) = p(t_1)$ and $u(t) = u(t_1)$ stay constant

General case

Theorem

The solution path to

$$\dot{p}_{+}(t) = A^{*}(b - Au(t)), \quad p(t) \in \partial ||u(t)||_{1}$$

with initial conditions $t_0 = 0$, p(0) = 0, u(0) = 0, is uniquely given by:

• for
$$k = 1, 2, ..., K$$

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- for $k = 1, 2, \ldots, K$
 - p(t) is piece-wise linear

$$p(t) = p(t_{k-1}) + (t - t_{k-1})A^*(b - Au(t_{k-1})), \quad t \in [t_{k-1}, t_k],$$

where

$$t_k := \sup\{t > t_{k-1} : p(t) \in ||u(t_{k-1})||_1\}.$$

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• $u(t) = u(t_{k-1})$ for $t \in [t_{k-1}, t_k)$; if $t_k \neq \infty$, compute

$$u(t_k) = \underset{u}{\arg\min} ||Au - b||_2^2 \quad \text{s.t. } u_i \begin{cases} \ge 0, \quad p_i(t_k) = 1, \\ = 0, \quad p_i(t_k) \in (-1, 1), \\ \le 0, \quad p_i(t_k) = -1. \end{cases}$$













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Differences:

- OMP evolves index set S; new method evolves l₁-subgradient p, keeping more information
- both add one nonzero each iteration, but new method may also drop
- both have extensions to have multiple adds/drops each iteration
- similar computing cost at each iteration

Numerical results: new method is more powerful than OMP at sparse recovery

Relation to LASSO

Model:

$$\min \|u\|_1 + \frac{t}{2n} \|Au - b\|_2^2$$

Optimality conditions:

$$\frac{p}{t} = A^*(b - Au), \quad p \in \partial ||u||_1.$$

Similarities:

- $p \in \partial \|u\|_1 \cap \mathcal{R}(A^T)$, and p is continuous
- as $t \to \infty$, both u is a solution to

$$\min \|u\|_1 \quad \text{s.t. } Au = b.$$

- as t increases, both add and can also drop predictors
- sign consistency under conditions

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Qualitative differences:

- · to reach the same fitting, new method requires fewer nonzeros
- given the same number of nonzeros, new method has better fitting
- LASSO is biased; new method is not

There are differences in both variable selection and estimation

Bias

Suppose both method select true $S = \operatorname{supp}(u^*)$.

LASSO gives

$$\hat{u}_S(\tau) = (A_S^T A_S)^{-1} A_S b - \underbrace{\frac{m}{\underline{\tau}} \operatorname{sign}(\hat{u}_S(\tau))}_{\text{bias}}$$

more noise \implies smaller $\tau \implies$ stronger bias

new method gives

$$u_S(t) = (A_S^T A_S)^{-1} A_S b$$

• assuming 0-mean noise, $u_S(t)$ is unbiased since

$$\mathbf{E}[u_S(t)] = \mathbf{E}[(A_S^T A_S)^{-1} A_S(A_S u_S^* + \epsilon)] = u_S^*$$



Theorem

- 1. For any A and b, solution to (1) exists;
- 2. p(t) is unique and piece-wise linear;
- 3. Au(t) b is piece-wise constant; ||Au(t) b|| is non-increasing;
- 4. There exists a piece-wise constant u(t);
- 5. Let I = supp(u(t)) and assume 0-mean noise. Then, u(t) is an unbiased solution to

$$A_I u_I = b_I$$

6. There exists t_{∞} such that for $t \ge t_{\infty}$, $u(t) = u_{\infty}$ is a solution to

$$\min \|u\|_1 \quad s.t. \|Au - b\|_2 = \min_w \|Aw - b\|_2.$$

Many results are essentially known from CAM 04-13 and 11-08.

Prostate tumor size

- select predictors among 8 clinical features to predict prostate tumor size
 - apply 4 different methods to 67 training cases

Predictor	LS	Subset	glmnet	ISS
Intercept	2.452	2.466	2.481	2.476
lcavol	0.716	0.667	0.622	0.554
lweight	0.293	0.366	0.289	0.279
age	-0.143	0	-0.096	0
lbph	0.212	0	0.188	0.198
svi	0.310	0.268	0.262	0.238
lcp	-0.289	-0.291	-0.164	0
gleason	-0.021	0	0	0
pgg45	0.277	0.227	0.187	0.122
Test Error	0.586	0.587	0.543	0.541

results were tested on 30 testing cases

LS = least squares, Subset = best subset regression glmnet = a package with LASSO, proposed approach

Cross validation



ISS achieves better fitting with fewer nonzerso than LASSO (glmnet) Note: exactly the same cross validation was applied to both methods

Relation to Bregman iteration

• Discretize
$$\dot{p} = A^*(b - Au)$$
 by

$$p^{k+1} = p^k + \delta A^* (b - Au^k).$$

It is the first-order optimality condition to Bregman iteration

$$\begin{split} u^{k+1} &\leftarrow \min D_{\|\cdot\|_1}(u; u^k) + \frac{\delta}{2n} \|Au - b\|^2, \\ \text{where} \quad D_{\|\cdot\|_1}(u; u^k) := \|u\|_1 - \|u^k\|_1 - \langle p^k, u - u^k \rangle. \end{split}$$

After change of variable (CAM 04-13, 07-37)

$$u^{k+1} \leftarrow \min \|u\|_1 + \frac{\delta}{2n} \|Au - b^k\|^2,$$

$$b^{k+1} \leftarrow b^k + (b - Au^k).$$

- Still true if $\|\cdot\|_1$ is replace by any convex regularizer

Sparse recovery from noisy measurements





LASSO

Bregman

• Damping
$$\dot{p} = A^*(b - Au)$$
 into

$$\dot{p}(t) + \alpha \dot{u}(t) = A^*(b - Au(t)).$$

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• Forward Euler discretization

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• Can be simplified to (in a miracle way!)

$$u^{k+1} = \alpha^{-1} \operatorname{shrink}(A^T y^k)$$
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- If b is noisy, stop at a finite k for best solution.
- If b is noise-free, u^k converges at a linear rate to the solution of

$$\min \|u\|_1 + \frac{\alpha}{2} \|u\|_2^2 \quad \text{s.t.} \ \|Au - b\|_2 = \min_w \|Aw - b\|_2.$$

Sufficiently small α (e.g., $\alpha < \frac{1}{10 \|u^*\|_{\infty}}$ in CS) $\implies u^*$ is an ℓ_1 minimizer

Apply different primal and dual algorithms to the same model

$$\min \|u\|_1 + \frac{t}{2n} \|Au - b\|_2^2.$$

Dual algorithms do better than the model!



Path consistency

Question: does there $\exists t$ so that solution u(t) has the following properties?

- no false positive: if $u_i = 0$, then $u_i(t) = 0$
- no false negative: if $u_i \neq 0$, then $u_i(t) \neq 0$
- sign consistency: furthermore, sign(u) = sign(u(t)).

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Theorem

Under the Assumptions

- Gaussian noise: $\omega \sim N(0, \sigma^2 I)$,
- normalized column: $\frac{1}{n} \max_j ||A_j||^2 \le 1$,

and under appropriate conditions, the new method has sign consistency and gives an unbias estimate to u^* .

Proof is based on the next two lemmas.

No false positive

Define true support $S := \operatorname{supp}(u)$, and let $T := S^c$.

Lemma

Under **Assumptions**, if A_S has full column rank and

$$\max_{j \in T} \|A_j^T A_S (A_S^T A_S)^{-1}\|_1 \le 1 - \eta$$

for some $\eta \in (0,1)$, then with high probability

$$\operatorname{supp}(u(s)) \subseteq S, \quad \forall s \leq \overline{t} := O\left(\frac{\eta}{\sigma}\sqrt{\frac{m}{\log n}}\right).$$

Proof uses: (i) concentration inequality and (ii) if $supp(u(s)) \subseteq S, s \leq t$, then

$$p(s)_T = A_T^T A_S (A_S^T A_S)^{-1} p(s)_S + t A_T^* P_{A_S^{\perp}} w, \quad s \le t.$$

No false negative / sign consistency

Lemma

Under Assumptions, if $A_S^*A_S \succeq \gamma I$ and

$$u_{\min} \ge \max\left\{O\left(\frac{\sigma}{\sqrt{\gamma}}\sqrt{\frac{\log|S|}{m}}\right), O\left(\frac{\sigma\log|S|}{\eta\gamma}\sqrt{\frac{\log n}{m}}\right)\right\},\$$

then there exist t^* (which can be given explicitly) so that with high probability

 $\operatorname{sign}(u(t)) = \operatorname{sign}(u)$

and $u(t) = u_S - (A_S^*A_S)^{-1}A_S^*\omega$ obeys

$$\|u(t) - u\|_{\infty} \le u_{\min}/2.$$

- first term in max ensures $\|(A_S^*A_S)^{-1}A_S^*\omega\|_{\infty} \leq u_{\min}/2$
- second term ensures: $\inf\{t : \operatorname{sign}(u_S(t)) = \operatorname{sign}(u_S)\} \leq \overline{t}.$